

Bernoulli 试验中 k 阶概率分布的众数

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摘要: 分别讨论了 k 阶负二项分布 $NB_k(r, p)$ 当参数 $p = 0.5$, $k = 2$ 时的众数及 k 阶 Poisson 分布 $P_k(\lambda)$ 在参数 λ 和 k 取某些特定值时的众数。同时提出一个关于 k 阶二项分布 $B_k(n, p)$ 众数的猜想。

关键词: 系统可靠性; k 阶分布; 众数; 概率母函数; 分布律; Fibonacci 数列

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On the modes of distributions of order k in Bernoulli trials

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Abstract: The modes of the negative binomial distribution of order k and the Poisson distribution of the same order are investigated for some fixed parameters. Furthermore, a conjecture on the mode of the binomial distribution of order k is offered.

Keywords: system reliability; distribution of order k ; mode; probability generating function; probability distribution function; Fibonacci sequence

在成功参数为 p ($0 < p < 1$) 的独立 Bernoulli 试验序列中, 以 $X_{k,r}$ 表示长度为 k 的成功游程^[1-7] 发生 r 次的试验次数(等待时间), 称随机变量 $X_{k,r}$ 服从参数为 (r, p) 的 k 阶负二项分布, 记作 $NB_k(r, p)$ ^[8-13]。关于该分布, 有下面的结论^[11,14]:

引理 1 设随机变量 $X_{k,r} \sim NB_k(r, p)$, 则 $X_{k,r}$ 的概率母函数为

$$G(X_{k,p}; x) = \left(\frac{p^k x^k (1 - px)}{1 - x + qp^k x^{k+1}} \right)^r.$$

另外, 从 k 阶负二项分布 $NB_k(r, p)$ 出发, 可以导出 k 阶 Poisson 分布 $P_k(\lambda)$ 的概率母函数及其概率分布律^[9,14]:

引理 2 设随机变量 $Z_k \sim P_k(\lambda)$, 则 Z_k 的概

率母函数为

$$G(Z_k; x) = e^{\lambda(x + x^2 + \dots + x^{k-1})},$$

Z_k 的概率分布律为

$$P(Z_k = m) = \sum_{\substack{(m_1, m_2, \dots, m_k) \\ m_1 + 2m_2 + \dots + km_k = m}} \frac{\lambda^{m_1 + m_2 + \dots + m_k}}{m_1! m_2! \dots m_k!} e^{-\lambda k}, m = 0, 1, \dots.$$

在参数为 p 的 n 重独立 Bernoulli 试验中, 以 $N_n^{(k)}$ 表示长度为 k ($\leq n$) 的成功游程发生的次数, 则称随机变量 $N_n^{(k)}$ 服从参数为 (n, p) 的 k 阶二项分布, 记作 $N_n^{(k)} \sim B_k(n, p)$ 。仍从 k 阶负二项分布 $NB_k(r, p)$ 出发, 可以辗转导出随机变量 $N_n^{(k)}$ 的概率分布律^[2] 和数学期望如下^[15]:

引理 3 设随机变量 $N_n^{(k)} \sim B_k(n, p)$, 则 $N_n^{(k)}$

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的概率分布律为

$$P(N_n^{(k)} = r) = \sum_{s=0}^{k-1} \sum_{\substack{(m_1, m_2, \dots, m_k) \\ m_1+2m_2+\dots+km_k \\ =n-s-kr}} \binom{m_1 + \dots + m_k + r}{m_1, m_2, \dots, m_k, r} \left(\frac{q}{p}\right)^{m_1+\dots+m_k} p^n,$$

$r=0,1,2,\dots,[n/k]$. 其中 $[x]$ 表示不超过 x 的最大整数.

引理4 设随机变量 $N_n^{(k)} \sim B_k(n, p)$, 则

$N_n^{(k)}$ 的数学期望为

$$EN_n^{(k)} = \sum_{m=0}^{[n/k]} \{1 + (n - mk)q\} p^{mk}.$$

我们知道, 众数(mode)^[16-18]是概率分布的一个重要统计量, 即概率最大时随机变量对应的取值. 它可广泛应用于抽样检查理论研究^[19-20]. 许多经典分布的众数已经为人熟知, 例如, 几何分布的众数是1, 参数 λ 为正整数时Poisson分布的众数是 λ 或 $(\lambda-1)$ 等等. 然而, 许多高阶分布诸如 k 阶负二项分布、 k 阶Poisson分布和 k 阶二项分布等等, 它们的众数尚属未知. 文献[16-18]讨论了某些特定参数下 k 阶负二项分布 $NB_k(r, p)$ 和 k 阶Poisson分布 $P_k(\lambda)$ 的众数问题. 延续上述文献, 本文作者对这两个分布的众数做进一步的研究, 此外, 对与它们密切相关的 k 阶二项分布 $B_k(n, p)$ 的众数, 也做了一些讨论.

1 k 阶负二项分布的众数

本节约定参数 $p=0.5$. 论证过程中使用了Fibonacci数列及其推广形式. 事实上, 它在概率分布理论当中有着广泛而重要的应用^[21-24].

定理1 当参数 $k=2, r=1$ 时, 随机变量 $X_{2,1} \sim NB_2(1, 0.5)$ 有唯一的众数 $m_{2,1}=2$.

证明 记随机变量 $X_{2,1}$ 的概率母函数为 $G_{2,1}(x)$, 则由引理1得到

$$\begin{aligned} G_{2,1}(x) &= \frac{0.5^2 x^2 (1 - 0.5x)}{1 - x + 0.5^3 x^3} = \frac{x^2}{2^2} \frac{4}{4 - 2x - x^2} = \\ &\quad \frac{x^2}{2^2} \frac{4}{(\sqrt{5} + 1 + x)(\sqrt{5} - 1 - x)} = \\ &\quad \frac{x^2}{2^2} \frac{2}{\sqrt{5}} \left(\frac{1}{\sqrt{5} + 1 + x} + \frac{1}{\sqrt{5} - 1 - x} \right) = \\ &\quad \frac{x^2}{2^2} \frac{2}{\sqrt{5}} \left[\frac{1}{\sqrt{5} + 1} \sum_{n=0}^{\infty} \left(-\frac{x}{\sqrt{5} + 1} \right)^n + \right. \\ &\quad \left. \frac{1}{\sqrt{5} - 1} \sum_{n=0}^{\infty} \left(\frac{x}{\sqrt{5} - 1} \right)^n \right] = \\ &\quad \frac{x^2}{2^2} \frac{2}{\sqrt{5}} \left[\frac{\sqrt{5} - 1}{4} \sum_{n=0}^{\infty} \left(\frac{1 - \sqrt{5}}{4} \right)^n x^n + \right. \\ &\quad \left. \frac{\sqrt{5} + 1}{4} \sum_{n=0}^{\infty} \left(\frac{1 + \sqrt{5}}{4} \right)^n x^n \right] = \end{aligned}$$

$$\begin{aligned} &\frac{1 + \sqrt{5}}{4} \sum_{n=0}^{\infty} \left(\frac{1 + \sqrt{5}}{4} \right)^n x^n = \\ &\frac{x^2}{2^2} \frac{1}{\sqrt{5}} \left[\frac{\sqrt{5} - 1}{2} \sum_{n=0}^{\infty} \left(\frac{1 - \sqrt{5}}{2} \right)^n \frac{x^n}{2^n} + \right. \\ &\quad \left. \frac{1 + \sqrt{5}}{2} \sum_{n=0}^{\infty} \left(\frac{1 + \sqrt{5}}{2} \right)^n \frac{x^n}{2^n} \right] = \\ &\frac{x^2}{2^2} \frac{1}{\sqrt{5}} \left[\sum_{n=0}^{\infty} \left(\frac{1 + \sqrt{5}}{2} \right)^{n+1} \frac{x^n}{2^n} - \right. \\ &\quad \left. \sum_{n=0}^{\infty} \left(\frac{1 - \sqrt{5}}{2} \right)^{n+1} \frac{x^n}{2^n} \right] = \\ &\frac{x^2}{2^2} \sum_{n=0}^{\infty} \frac{1}{\sqrt{5}} \left[\left(\frac{1 + \sqrt{5}}{2} \right)^{n+1} - \left(\frac{1 - \sqrt{5}}{2} \right)^{n+1} \right] \frac{x^n}{2^n} = \\ &\frac{x^2}{2^2} \sum_{n=0}^{\infty} \frac{F_n}{2^n} x^n \end{aligned} \quad (1)$$

其中 $\{F_n\}_{0}^{\infty}$ 为著名的Fibonacci数列, 满足: $F_0 = F_1 = 0$, $F_n = F_{n-1} + F_{n-2}$, $n \geq 2$. 由此得到随机变量 $X_{2,1}$ 的概率分布律

$$P_{n+2} = P(X_{2,1} = n+2) = F_n / 2^{n+2}, n = 0, 1, \dots.$$

考虑到 $F_1 < 2F_0$, $F_2 = 2F_1$, 且当 $n \geq 2$ 时有 $F_{n+1} = F_n + F_{n-1} < 2F_n$, 结合 $P_{n+2} - P_{n+3} = (2F_n - F_{n+1}) / 2^{n+3}$, 因此 $P_2 > P_3$, $P_3 = P_4 > P_5 > P_6 > \dots$. 从而有 $P_2 = \max\{P_{n+2}, n \geq 0\}$, 所以 $m_{2,1} = 2$. 证毕.

为了讨论参数 $k=2, r \geq 2$ 时随机变量 $X_{2,r} \sim NB_2(r, 0.5)$ 的众数问题, 我们记

$$\begin{aligned} F_n^{(1)} &= F_n, \\ F_n^{(2)} &= F_0 F_n + F_1 F_{n-1} + \dots + F_n F_0, \\ F_n^{(3)} &= F_0 F_n^{(2)} + F_1 F_{n-1}^{(2)} + \dots + F_n F_0^{(2)}, \dots, \\ F_n^{(r)} &= F_0 F_n^{(r-1)} + F_1 F_{n-1}^{(r-1)} + \dots + F_n F_0^{(r-1)}. \end{aligned}$$

称 $\{F_n^{(r)}\}_{0}^{\infty}$ 为 r 阶Fibonacci数列, $r=1, 2, 3, \dots$. 则有下面的性质:

$$\begin{aligned} F_n^{(1)} &\equiv F_n; F_0^{(r)} = 1; F_1^{(r)} = r; \\ F_2^{(r)} &= r(r+3)/2; \\ F_{n+1}^{(r)} &= F_n^{(r)} + F_{n-1}^{(r)} + F_{n+1}^{(r-1)}, r \geq 2. \end{aligned}$$

定理2 当参数 $k=2$ 时, 随机变量 $X_{2,r} \sim NB_2(r, 0.5)$ 的概率母函数为

$$G_{2,r}(x) = \frac{x^{2r}}{2^{2r}} \sum_{n=0}^{\infty} \frac{F_n^{(r)}}{2^n} x^n, r = 1, 2, \dots.$$

证明 (归纳法) 当 $r=1$ 时结论显然. 设 $(r-1)$ 时成立, 即有

$$G_{2,r-1}(x) = \frac{x^{2(r-1)}}{2^{2(r-1)}} \sum_{n=0}^{\infty} \frac{F_n^{(r-1)}}{2^n} x^n.$$

结合引理1和式(1), 得到

$$G_{2,r}(x) = G_{2,r-1}(x) G_{2,1}(x) = \frac{x^{2(r-1)}}{2^{2(r-1)}} \sum_{n=0}^{\infty} \frac{F_n^{(r-1)}}{2^n} x^n.$$

$$\begin{aligned} \frac{x^2}{2^2} \sum_{n=0}^{\infty} \frac{F_n}{2^n} x^n &= \frac{x^{2r}}{2^{2r}} \sum_{n=0}^{\infty} \frac{F_n^{(r-1)}}{2^n} x^n \cdot \sum_{n=0}^{\infty} \frac{F_n}{2^n} x^n = \\ \frac{x^{2r}}{2^{2r}} \sum_{n=0}^{\infty} \frac{F_0 F_n^{(r-1)} + F_1 F_n^{(r-1)} + \dots + F_n F_0^{(r-1)}}{2^n} x^n &= \\ \frac{x^{2r}}{2^{2r}} \sum_{n=0}^{\infty} \frac{F_n^{(r)}}{2^n} x^n. \end{aligned}$$

定理得证.

定理 3 当参数 $k=2, r=2$ 时, 随机变量 $X_{2,2} \sim NB_2(2,0.5)$ 的众数为 $m_{2,2}=6,7,8$.

证明 记随机变量 $X_{2,2}$ 的概率分布律为 $P_n = P(X_{2,2}=n)$, 由定理 2 得 $P_{n+4} = F_n^{(2)}/2^{n+4}$, 通过计算 2 阶 Fibonacci 数列 $\{F_n^{(2)}\}_{n=0}^{\infty}$ 各项值得 $P_4 = F_0^{(2)}/2^4 = 1/2^4, P_5 = F_1^{(2)}/2^5 = 2/2^5, P_6 = F_2^{(2)}/2^6 = 5/2^6, P_7 = F_3^{(2)}/2^7 = 10/2^7, P_8 = F_4^{(2)}/2^8 = 20/2^8$, 因此有 $P_4 = P_5 < P_6 = P_7 = P_8$.

当项数 $n+4 \geq 9$ 时, 利用 2 阶 Fibonacci 数列的性质, 我们有

$$\begin{aligned} P_{n+4} - P_{n+3} &= F_n^{(2)}/2^{n+4} - F_{n-1}^{(2)}/2^{n+3} = \\ [F_n^{(2)} - 2F_{n-1}^{(2)}]/2^{n+4} &= [F_n + F_{n-2}^{(2)} - F_{n-1}^{(2)}]/2^{n+4} = \\ [F_n - F_{n-1} - F_{n-3}^{(2)}]/2^{n+4} &= [F_{n-2} - F_{n-3}^{(2)}]/2^{n+4} = \\ [F_{n-3} - F_{n-4} - F_{n-3}^{(2)}]/2^{n+4} &< 0. \end{aligned}$$

这意味着 $P_8 > P_9 > P_{10} > \dots$. 综上述, 有 $P_6 = P_7 = P_8 = \max\{P_{n+4}, n \geq 0\}$. 证毕.

定理 4 当参数 $k=2, r=3$ 时, 随机变量 $X_{2,3} \sim NB_2(3,0.5)$ 的众数为 $m_{2,3}=13$.

证明 记随机变量 $X_{2,3}$ 的概率分布律为 $P_n = P(X_{2,3}=n)$, 由定理 2 得 $P_{n+6} = F_n^{(3)}/2^{n+6}$, 通过计算 3 阶 Fibonacci 数列 $\{F_n^{(3)}\}_{n=0}^{\infty}$ 各项值得 $P_6 = F_0^{(3)}/2^6 = 1/2^6, P_7 = F_1^{(3)}/2^7 = 3/2^7, P_8 = F_2^{(3)}/2^8 = 9/2^8, P_9 = F_3^{(3)}/2^9 = 22/2^9, P_{10} = F_4^{(3)}/2^{10} = 51/2^{10}, P_{11} = F_5^{(3)}/2^{11} = 111/2^{11}, P_{12} = F_6^{(3)}/2^{12} = 233/2^{12}, P_{13} = F_7^{(3)}/2^{13} = 474/2^{13}, P_{14} = F_8^{(3)}/2^{14} = 942/2^{14}, P_{15} = F_9^{(3)}/2^{15} = 1836/2^{15}$. 因此有 $P_6 < P_7 < P_8 < P_9 < P_{10} < P_{11} < P_{12} < P_{13}$ 以及 $P_{13} > P_{14} > P_{15}$.

记 $\delta_n = P_n - P_{n-1}$ 且假设 $\delta_n < 0, \delta_{n+1} < 0$, 则由

$$\begin{aligned} \delta_n &= P_n - P_{n-1} = F_{n-6}^{(3)}/2^n - F_{n-7}^{(3)}/2^{n-1} = \\ [F_{n-6}^{(3)} - 2F_{n-7}^{(3)}]/2^n &= \\ [F_{n-6}^{(3)} + F_{n-8}^{(3)} - F_{n-7}^{(3)}]/2^n &= \\ [F_{n-6}^{(3)} - F_{n-7}^{(3)} - F_{n-9}^{(3)}]/2^n &= \\ [F_{n-8}^{(3)} + F_{n-6}^{(3)} - F_{n-9}^{(3)}]/2^n &< 0 \end{aligned}$$

得到

$$2^n \delta_n = F_{n-8}^{(3)} + F_{n-6}^{(3)} - F_{n-9}^{(3)} < 0$$

以及

$$2^{n+1} \delta_{n+1} = F_{n-7}^{(3)} + F_{n-5}^{(3)} - F_{n-8}^{(3)} < 0$$

则

$$\begin{aligned} 2^{n+2} \delta_{n+2} &= F_{n-6}^{(2)} + F_{n-4}^{(3)} - F_{n-7}^{(3)} = \\ (F_{n-7}^{(2)} + F_{n-8}^{(2)} + F_{n-6}^{(3)}) &+ (F_{n-5}^{(3)} + F_{n-6}^{(3)}) - \\ (F_{n-8}^{(3)} + F_{n-9}^{(3)} + F_{n-7}^{(2)}) &= \\ (F_{n-8}^{(2)} + F_{n-6}^{(3)} - F_{n-9}^{(3)}) &+ F_{n-5}^{(3)} + F_{n-6}^{(3)} - F_{n-8}^{(3)} = \\ 2^n \delta_n + F_{n-5}^{(3)} + F_{n-6}^{(3)} - F_{n-8}^{(3)} &= \\ 2^n \delta_n + (F_{n-5}^{(2)} + F_{n-7}^{(2)} - F_{n-8}^{(3)}) &+ F_{n-6}^{(3)} - F_{n-7}^{(2)} = \\ 2^n \delta_n + 2^{n+1} \delta_{n+1} + (F_{n-7}^{(2)} + F_{n-8}^{(3)} - F_{n-7}^{(2)}) &< 0. \end{aligned}$$

因此有 $\delta_{n+2} < 0$.

这表明, 只要有连续两个 δ_n 小于零, 则后面的 δ_n 均小于零. 由 $P_{13} > P_{14} > P_{15}$ 得到 $\delta_{14} < 0$ 和 $\delta_{15} < 0$, 因此 $\delta_n < 0, n=16, 17, 18, \dots$. 即有 $P_{15} > P_{16} > P_{17} > P_{18} > \dots$. 所以 $P_{13} = \max\{P_{n+6}, n \geq 0\}$. 证毕.

注记 通过更复杂的推导论证可得到随机变量 $X_{2,4}$ 的众数 $m_{2,4}=19$; $X_{2,5}$ 的众数 $m_{2,5}=25$; $X_{2,6}$ 的众数 $m_{2,6}=31$; $X_{2,7}$ 的众数 $m_{2,7}=37$; $X_{2,8}$ 的众数 $m_{2,8}=43$; $X_{2,9}$ 的众数 $m_{2,9}=49$; $X_{2,10}$ 的众数 $m_{2,10}=55$. 由此可得, 随机变量 $X_{2,r} \sim NB_2(r, 0.5)$ 的众数 $m_{2,r}=6r-5, r \geq 3$.

2 k 阶 Poisson 分布的众数

约定参数 $k=2$, 展开本节的讨论.

定理 5 当参数 $0 < \lambda < \sqrt{3}-1$ 时, 随机变量 $Z_2 \sim P_2(\lambda)$ 的众数为 $m_{Z_2}=0$; 当参数 $\lambda=\sqrt{3}-1$ 时, 众数为 $m_{Z_2}=0$ 或 2; 当参数 $\sqrt{3}-1 < \lambda \leq 1$ 时, 众数为 $m_{Z_2}=2$.

证明 由引理 2 得

$$P_m = P(Z_2=m) = \sum_{\substack{(m_1, m_2) \\ m_1+2m_2=m}} \frac{\lambda^{m_1+m_2}}{m_1! m_2!} e^{-2\lambda},$$

所以有 $P_0 = e^{-2\lambda}$, $P_1 = \lambda e^{-2\lambda}$, $P_2 = (\lambda + \lambda^2/2) e^{-2\lambda}$. 注意到 $0 < \lambda \leq 1$, 当 $n \geq 1$ 时, 有

$$\begin{aligned} P_{2n} - P_{2n+1} &= \\ \sum_{s=0}^n \frac{\lambda^{2n-s} e^{-2\lambda}}{(2n-2s)! s!} - \sum_{s=0}^n \frac{\lambda^{2n+1-s} e^{-2\lambda}}{(2n+1-2s)! s!} &\geq \\ \sum_{s=0}^n \frac{\lambda^{2n-s} e^{-2\lambda}}{(2n-2s)! s!} - \sum_{s=0}^n \frac{\lambda^{2n-s} e^{-2\lambda}}{(2n+1-2s)! s!} &= \\ \sum_{s=0}^n \left[\frac{\lambda^{2n-s} e^{-2\lambda}}{(2n-2s)! s!} - \frac{\lambda^{2n-s} e^{-2\lambda}}{(2n+1-2s)! s!} \right] &> 0, \\ P_{2n+1} - P_{2n+2} &= \\ \sum_{s=0}^n \frac{\lambda^{2n+1-s} e^{-2\lambda}}{(2n+1-2s)! s!} - \sum_{s=0}^{n+1} \frac{\lambda^{2n+2-s} e^{-2\lambda}}{(2n+2-2s)! s!} &= \\ \sum_{s=0}^{n-1} \left[\frac{\lambda^{2n+1-s} e^{-2\lambda}}{(2n+1-2s)! s!} - \frac{\lambda^{2n+2-s} e^{-2\lambda}}{(2n+2-2s)! s!} \right] &+ \end{aligned}$$

$$e^{-2\lambda} \left(\frac{\lambda^{n+1}}{1! n!} - \frac{\lambda^{n+2}}{2! n!} - \frac{\lambda^{n+1}}{0! (n+1)!} \right) > 0.$$

因此当 $0 < \lambda \leq 1$ 时有 $P_2 > P_3 > P_4 > \dots$. 综上所述, 有结论:

当 $\lambda = \sqrt{3} - 1$ 时, $P_0 = P_2 = \max\{P_m, m \geq 0\}$,

即得众数 $m_{Z_2} = 0$ 或 2; 当 $0 < \lambda < \sqrt{3} - 1$ 时, $P_0 = \max\{P_m, m \geq 0\}$, 即得众数 $m_{Z_2} = 0$; 当 $\sqrt{3} - 1 < \lambda \leq 1$ 时, $P_2 = \max\{P_m, m \geq 0\}$, 即得众数 $m_{Z_2} = 2$. 证毕.

定理6 当参数 $\lambda > 1$ 且 $\lambda \in N$ 时, 随机变量 $Z_2 \sim P_2(\lambda)$ 的众数为 $m_{Z_2} = 3\lambda - 1$.

证明 由引理2得随机变量 $Z_2 \sim P_2(\lambda)$ 的概率母函数为

$$G(x) = G(Z_2; x) = e^{\lambda(x+x^2-2)},$$

对其求导, 得

$$G'(x) = \lambda(1+2x)G(x),$$

将上述函数两边同时求 $(n-1)$ 阶导数并令 $x=0$ 得

$$G^{(n)}(0) = \lambda G^{(n-1)}(0) + 2\lambda(n-1)G^{(n-2)}(0).$$

结合概率母函数 $G(x)$ 与概率分布律 P_n 之间的关系 $P_n = G^{(n)}(0)/n!$ 可得

$$nP_n = \lambda(P_{n-1} + 2P_{n-2}) \quad (2)$$

假定随机变量 Z_2 的众数是 $m_{Z_2} = n^*$, 那么由式(2)得

$$n^* P_{n^*} = \lambda(P_{n^*-1} + 2P_{n^*-2}) \leq 3\lambda P_{n^*},$$

得到

$$n^* \leq 3\lambda \quad (3)$$

令 $\delta_0 = P_0 > 0$, $\delta_n = P_n - P_{n-1}$, $n \geq 1$. 则有

$$\delta_1 = P_1 - P_0 = (\lambda - 1)e^{-2\lambda} > 0,$$

$$\delta_2 = P_2 - P_1 = \lambda^2 e^{-2\lambda}/2 > 0 \quad (4)$$

由式(2)得

$$\delta_{n+2} = \frac{\lambda(\lambda+n)\delta_n}{(n+1)(n+2)} + \frac{\lambda(3\lambda-n-4)}{(n+1)(n+2)}P_{n-1} \quad (5)$$

因此, 当 $1 \leq n \leq 3\lambda - 4$ 时, 可得 $\delta_{n+2} > 0$, 即

$$\delta_3 > 0, \delta_4 > 0, \delta_5 > 0, \dots, \delta_{3\lambda-1} > 0 \quad (6)$$

结合式(4)和(6)得到 $P_0 < P_1 < P_2 < \dots < P_{3\lambda-2}$, 从而可知

$$3\lambda - 2 \leq n^* \quad (7)$$

将式(3)和(7)联立, 得

$$3\lambda - 2 \leq n^* \leq 3\lambda \quad (8)$$

由式(2)可得

$$n\delta_n = (\lambda - n)\delta_{n-1} + (3\lambda - n)P_{n-2} \quad (9)$$

令 $n = 3\lambda$ 代入上式(9)得

$$\delta_{3\lambda} = -2\delta_{3\lambda-1} \quad (10)$$

令 $n = 3\lambda - 3, 3\lambda - 5, 3\lambda - 7$ 代入式(9)得

$$\begin{cases} \delta_{3\lambda-1} = \frac{\lambda(4\lambda-3)\delta_{3\lambda-3} - \lambda P_{3\lambda-4}}{(3\lambda-2)(3\lambda-1)} \\ \delta_{3\lambda-3} = \frac{\lambda(4\lambda-5)\delta_{3\lambda-5} + \lambda P_{3\lambda-6}}{(3\lambda-4)(3\lambda-3)} \\ \delta_{3\lambda-5} = \frac{\lambda(4\lambda-7)\delta_{3\lambda-7} + 3\lambda P_{3\lambda-8}}{(3\lambda-6)(3\lambda-5)} \end{cases} \quad (11)$$

仍由式(2)可得

$$\begin{cases} P_{3\lambda-4} = \frac{\lambda(7\lambda-10)P_{3\lambda-6} + 2\lambda^2 P_{3\lambda-7}}{(3\lambda-4)(3\lambda-5)} \\ P_{3\lambda-6} = \frac{\lambda P_{3\lambda-7} + 2\lambda P_{3\lambda-8}}{3\lambda-6} \end{cases} \quad (12)$$

结合式(11)和(12), 得到

$$\begin{aligned} & \frac{(3\lambda-1)\cdots(3\lambda-6)}{\lambda^3} \delta_{3\lambda-1} = \\ & (64\lambda^3 - 267\lambda^2 + 360\lambda - 156)\delta_{3\lambda-7} + \\ & (3\lambda^2 + 24\lambda - 36)P_{3\lambda-8} > 0 \end{aligned} \quad (13)$$

结合式(10)及(13), 可以得到 $\delta_{3\lambda-1} < 0$ 和 $\delta_{3\lambda} > 0$, 这意味着

$$P_{3\lambda-1} > P_{3\lambda-2} \text{ 且 } P_{3\lambda-1} > P_{3\lambda}.$$

结合式(8)可得 Z_2 的众数是 $m_{Z_2} = n^* = 3\lambda - 1$. 证毕.

3 k阶二项分布的众数

仅在约定参数 $p=0.5, k=2, n=2n_0$ 情形下对随机变量 $N_{2n_0}^{(2)} \sim B_2(2n_0, 0.5)$ 进行讨论. 此时由引理4, 随机变量 $N_{2n_0}^{(2)}$ 的数学期望是

$$EN_{2n_0}^{(2)} = \sum_{m=0}^{\lfloor 2n_0/2 \rfloor} \{1 + (2n_0 - 2m)0.5\}0.5^{2m} = \frac{n_0}{3} - \frac{1 - 1/4n_0}{9},$$

我们给出猜想: $N_{2n_0}^{(2)}$ 的众数 $m_{N_{2n_0}^{(2)}} = [(n_0 - 1)/3]$.

下面仅举一例, 对其进行验证.

取 $n = 2n_0 = 10$, 由引理3得

$$\begin{aligned} P(N_{10}^{(2)} = 0) &= \frac{1}{2^{10}} \binom{10}{10, 0, 0} + \frac{1}{2^{10}} \binom{9}{8, 1, 0} + \dots \\ &+ \frac{1}{2^{10}} \binom{5}{0, 5, 0} + \frac{1}{2^{10}} \binom{9}{9, 0, 0} + \dots \\ &+ \frac{1}{2^{10}} \binom{8}{7, 1, 0} + \dots + \frac{1}{2^{10}} \binom{5}{1, 4, 0} = \frac{144}{1024}, \\ P(N_{10}^{(2)} = 1) &= \frac{1}{2^{10}} \binom{9}{8, 0, 1} + \frac{1}{2^{10}} \binom{8}{6, 1, 1} + \dots \\ &+ \frac{1}{2^{10}} \binom{5}{0, 4, 1} + \frac{1}{2^{10}} \binom{8}{7, 0, 1} + \frac{1}{2^{10}} \binom{7}{5, 1, 1} + \dots \\ &+ \frac{1}{2^{10}} \binom{6}{3, 2, 1} + \frac{1}{2^{10}} \binom{5}{1, 3, 1} = \frac{365}{1024}, \\ P(N_{10}^{(2)} = 2) &= \frac{1}{2^{10}} \binom{8}{6, 0, 2} + \frac{1}{2^{10}} \binom{7}{4, 1, 2} + \dots \end{aligned}$$

$$\begin{aligned} & \frac{1}{2^{10}} \binom{6}{2,2,2} + \frac{1}{2^{10}} \binom{5}{0,3,2} + \frac{1}{2^{10}} \binom{7}{5,0,2} + \\ & \frac{1}{2^{10}} \binom{6}{3,1,2} + \frac{1}{2^{10}} \binom{5}{1,2,2} = \frac{344}{1024}, \\ P(N_{10}^{(2)}=3) &= \frac{1}{2^{10}} \binom{7}{4,0,3} + \frac{1}{2^{10}} \binom{6}{2,1,3} + \\ & \frac{1}{2^{10}} \binom{5}{0,2,3} + \frac{1}{2^{10}} \binom{6}{3,0,3} + \frac{1}{2^{10}} \binom{5}{1,1,3} = \frac{145}{1024}, \\ P(N_{10}^{(2)}=4) &= \frac{1}{2^{10}} \binom{6}{2,0,4} + \frac{1}{2^{10}} \binom{5}{0,1,4} + \\ & \frac{1}{2^{10}} \binom{5}{1,0,4} = \frac{25}{1024}, \\ P(N_{10}^{(2)}=5) &= \frac{1}{2^{10}} \binom{5}{0,0,5} = \frac{1}{1024}, \end{aligned}$$

容易看出 $P(N_{10}^{(2)}=1)=\max_{0 \leq r \leq 5}\{P(N_{10}^{(2)}=r)\}$,

即有 $m_{N_{10}^{(2)}}=1$, 而由猜想众数公式可得 $m_{N_{10}^{(2)}}=[(n_0-1)/3]=[5-1/3]=1$, 与实际结果一致.

4 结论

本文讨论了独立 Bernoulli 试验序列中的三个 k 阶分布的众数问题:

1) k 阶负二项分布 $NB_k(r, p)$ 在参数 $p=0.5$, $k=2$ 时有 $m_{2,1}=2$; $m_{2,2}=6,7,8$; $m_{2,3}=13$, 一般地有 $m_{2,r}=6r-5, r \geq 3$.

2) k 阶 Poisson 分布 $P_k(\lambda)$ 当 $k=2, 0 < \lambda < \sqrt{3}-1$ 时, 有 $m_{Z_2}=0$; 当 $k=2, \lambda=\sqrt{3}-1$ 时, 有 $m_{Z_2}=0$ 或 2; 当 $k=2, \sqrt{3}-1 < \lambda \leq 1$ 时, 有 $m_{Z_2}=2$; 当 $k=2, \lambda > 1$ 且 $\lambda \in N$ 时, 有 $m_{Z_2}=3\lambda-1$.

3) k 阶二项分布 $B_k(n, p)$ 当参数 $p=0.5, k=2, n=2n_0$ 时有猜想 $m_{N_{2n_0}^{(2)}}=[(n_0-1)/3]$.

关于高阶分布众数的研究是概率统计理论的新领域, 其中的诸多问题包括本文提出的猜想, 有待更进一步探讨.

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